

# Systematic review and meta-analysis of maternal secondhand smoke exposure and neonatal outcomes

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## ABSTRACT

**Objectives:** To determine the risk of adverse pregnancy and foetal outcomes among nonsmoking women exposed to secondhand smoke (SHS) while pregnant and compare the quality of studies in middle- and high-income countries (MIC and HIC).

**Material and methods:** To assess the effect of maternal SHS exposure on preterm birth, low birth weight, stillbirth, and orofacial clefts, the authors searched three databases (2004–2015) and applied strict inclusion/exclusion criteria. Crude odds ratios (ORs), adjusted ORs, associated 95% confidence intervals, and raw data were used to estimate pooled ORs using random-effects models. Models were stratified by adjusted/unadjusted data, high study quality vs. low quality, and country income level.

**Results:** The authors identified 39 studies that assessed the effect of maternal SHS exposure on the four outcomes. Of these studies, the 20 conducted in MICs were of lower quality on average than those conducted in HICs. Evidence of a statistically significant effect of SHS exposure was found using adjusted pooled ORs among the high-quality studies for low birth weight (pOR: 1.282 [95% CI: 1.103–1.490]), preterm birth (pOR: 1.290 [95% CI: 1.127–1.476]), and stillbirth (pOR: 1.230 [95% CI: 1.037–1.459]). Using the high-quality adjusted ORs, the effect of SHS exposure in nonsmoking pregnant women on cleft palate with or without cleft lip did not reach significance (pOR: 1.307 [95% CI: 0.893–1.9]).

**Conclusions:** In all analyses, the odds of having negative outcomes were greater among studies conducted in MICs than in HICs. More research is needed in middle-income countries to inform policy and practice.

**KEY WORDS:** meta-analysis, second-hand smoke, reproductive health, pregnancy.

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## INTRODUCTION

A large body of evidence demonstrates that cigarette smoking and secondhand smoke (SHS) exposure during pregnancy pose serious risks to the mother and her developing child, including sudden infant death syndrome, low birth weight (LBW), orofacial defects (cleft lip with/without cleft palate [CLP]), preterm birth (PTB), and stillbirth [1]. While attention has been paid to reducing cigarette smoking during pregnancy, especially in high-income

countries (HICs), less progress has been made in reducing secondhand smoke exposure during pregnancy, especially in low- and middle-income countries (LMICs).

Substantial advances have been made over the past decade in implementing policies to reduce SHS exposure. Since 2007, 35 LMICs (25% of all LMICs) have put in place comprehensive smoke-free laws. Overall, comprehensive smoke-free legislation is currently in place in 55 countries covering 20% of the world's population

[2]. However, much of the world's population remains unprotected, and in countries where laws do exist, implementation and enforcement remain a challenge. Moreover, smoke-free policies generally target workplaces and public spaces, but not homes, so nonsmokers who live with a smoker remain affected. The World Health Organisation (WHO) estimates that 1.5 billion people are exposed to SHS at home, and 1.2 billion people are exposed to SHS in public places and workplaces [3].

In many LMICs, smoking rates among women remain low compared with men's rates. Thus, SHS exposure may be as important or more important to maternal and child health outcomes than active maternal smoking in these countries [4]. Of LMICs that have published rates of SHS exposure in women at home, rates are particularly high in China (63.9%), Indonesia (75.4%), Egypt (73.6%), and Vietnam (69.2%) [5]. In comparison, Caleyachetty *et al.*, using nationally representative household Demographic and Health Surveys in 54 countries, estimated that overall prevalence of any form of tobacco use in pregnant women was 2.6% [6].

Six previous meta-analyses have been conducted looking at SHS exposure and maternal and child health outcomes. Four included LBW [7-10], two included PTB [9, 10], one included CLP [11], and one included stillbirth [12]. An additional meta-analysis also analysed all orofacial clefts but did not conduct separate analyses for cleft lip [12]. A summary narrative of previous meta-analyses was published in 2013, which found, in general, a significant association between SHS exposure and LBW, PTB, and congenital malformations [13]. All of the mentioned meta-analyses found significantly increased odds of these health outcomes, with the exception of Salmasi *et al.* [10]. For LBW, PTB, and stillbirth, the most recent meta-analyses used 2009 data, and for orofacial clefts the most recent meta-analysis used 2013 data. However, additional studies have since been published, and previous analyses were limited primarily to studies conducted in HICs. For the purpose of this review, we evaluated studies focusing on the maternal and child health outcomes of low birth weight, preterm birth, orofacial cleft with or without cleft palate, and stillbirth, particularly seeking to include studies from LMICs.

## MATERIAL AND METHODS

### LITERATURE SEARCH STRATEGY – STUDY SELECTION

Two reviewers (J.B. and A.H.) used separate, comprehensive strategies to search the PubMed, Embase, and Web of Science (Appendix 1) databases for publications relevant to smoking and risk of adverse pregnancy outcomes published between 2004 and December 2015. The reviewers independently screened the titles and abstracts of all citations identified in the search. The full-text article was retrieved if either reviewer considered the cita-

tion potentially relevant. Disagreements or uncertainties were resolved by a third reviewer (M.P.).

### DATA ABSTRACTION

Two reviewers extracted data on author, title, outcome reviewed, study design, case/control or event/no-event data, crude odds ratios (ORs) and adjusted ORs (where presented), confounders, participant inclusion criteria, study sampling technique, and years of data included in the study. Differences in item coding were resolved by discussion.

### INCLUSION AND EXCLUSION CRITERIA

Case-control, cohort, and cross-sectional studies were included if they reported on LBW, PTB, cleft lip/palate, or stillbirth, as defined below, and compared mothers who were exposed to secondhand smoke with mothers who were not exposed, excluding mothers who were smokers. SHS exposure was determined either through self-report or via biochemical assay. Exposure could be domestic, occupational, or from other sources.

Studies were excluded if they: 1) were duplicates; 2) did not have at least English abstracts; or 3) did not account for maternal smoking or did not make clear whether active maternal smoking was a confounding factor. Experimental studies such as smoking cessation randomised control trials were also excluded.

### DEFINITION OF OUTCOMES

Four main outcomes, which the 2014 Surgeon General's report [1] reported to be causally linked with smoking and/or SHS exposure, were chosen as the focus of this paper:

- Low birth weight: A weight of less than 2500 g or 5 pounds 8 ounces at birth. Gestational age was not included in this definition because it was not always reported, particularly for lower quality studies.
- Preterm birth: A gestational age at birth of less than 37 weeks.
- Orofacial cleft (cleft lip with or without cleft palate): A clinical diagnosis of orofacial cleft lip with or without cleft palate, excluding any participants with a chromosomal abnormality or syndrome.
- Stillbirth: Birth of an infant that died in the womb after 20 weeks of pregnancy as determined by ultrasound, or ultrasound plus last menstrual period, and/or a minimum weight for gestational age (i.e. 400 g at 24 weeks).

### DEFINITION OF EXPOSURES

Studies that separated the mother's SHS exposure from the mother's active smoking were eligible for inclusion. Studies were included if they described the SHS exposure, which could be defined by biochemical validation or measuring air quality, or by mothers' self-report of: whether they were exposed (yes or no), the amount of smoke exposure in hours or minutes per day, the number of cigarettes per day that were smoked around them, or the source of

exposure (i.e. paternal, spouse, other). For biochemical validation, serum cotinine at acceptable levels was used to determine nonsmokers not exposed to SHS (under 0.01 ng/ml). For exposures where odds ratios were calculated for different amounts of exposure, the odds ratios were used for the maximum exposure time. For exposures that measured time in pregnancy (early vs. late), early pregnancy odds ratios were used. For both exposures, the ORs were chosen to increase the chances of seeing an effect of SHS on maternal and child outcomes.

### QUALITY ASSESSMENT

Study quality was assessed using an adapted Newcastle-Ottawa Scale – a nine-point scale assessing exposed (cases) and unexposed (controls) selection, confounding variables, measurement of secondhand smoke, and nonresponse rates. Two reviewers coded all eligible studies for quality.

### ANALYSIS

Tabulated data, crude estimates, and/or adjusted estimates were extracted from the included studies. A random effects model was chosen to account for heterogeneity of study populations and designs. Multiple separate analyses were conducted for each outcome: unadjusted full sample using either raw data or crude ORs, unadjusted high-quality sample using only studies that had a quality rating over 6, adjusted full sample using aORs to account for confounders, and adjusted high-quality studies. Additionally, studies were stratified by country income group (HIC vs. MIC) for each outcome. Analy-

ses included adjusted odds ratios to account for potential confounders, despite individual studies adjusting for different factors. Pooled odds ratios (pORs) and 95% confidence intervals were calculated for each analysis.

Heterogeneity between studies was assessed using the  $I^2$  statistic, which represents the percentage of total variation that is attributable to between-studies heterogeneity. Statistical significance of heterogeneity was analysed with the Q statistic. If  $I^2$  exceeded 75%, separate sensitivity analyses were conducted to determine the cause of this high heterogeneity. Data were analysed using the comprehensive meta-analysis software. Probability values below 0.05 were considered statistically significant.

## RESULTS

### LITERATURE SEARCH

Figure 1 shows the article-screening process. Of the identified publications, 5470 unique publications remained after duplicates were removed. Title and abstract screening resulted in exclusion of 4947 records, primarily due to irrelevance. Studies were then sorted by topic-LBW, PTB, CLP, and stillbirth, resulting in 198 remaining records. Finally, based on full text analysis, studies not meeting the exclusion criteria were removed, resulting in 39 studies included.

### STUDY CHARACTERISTICS AND QUALITY ASSESSMENT

Table 1 summarises the included studies [14-52]. An additional 13 studies were identified for LBW [14-25],

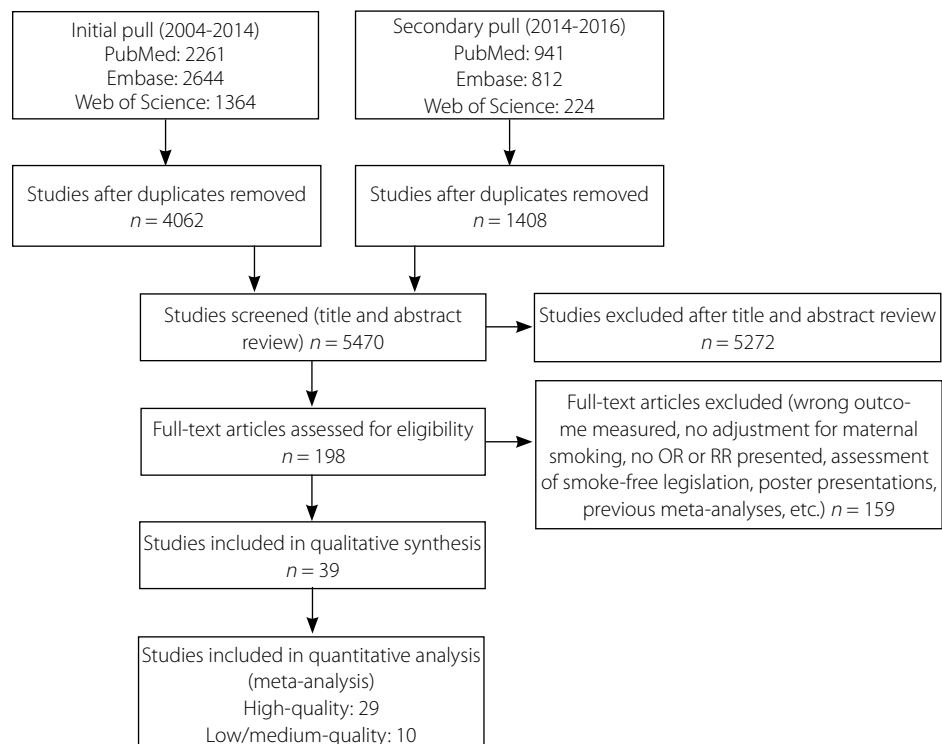


FIG. 1. Prisma diagram of study selection

TABLE 1. Information on included studies

First author, year	Outcome	Study design	No. of cases/ study size	Exposure	Risk estimates (aOR where applicable)	Adjusted factors	Income level, country	QA
Abu-Baker, 2010 [14]	LBW	Cohort	300	Home (y/n)	1.08 (1.03-1.12)	Mother's weight and height, weight gain, family income. Infant's gestational age	UM, Jordan	7
Abusalah, 2012 [15]	LBW	Case-control	446	Home or work (y/n)	3.4 (2.2-5.2)*	-	LM, Gaza	8
Andriani, 2014 [16]	LBW PTB	Cohort	3192 3192	Mean number of cigarettes smoked by spouse	0.89 (0.51-1.54) 1.16 (0.78-1.71)	Birth order, maternal age, parental BMI, and residence Birth order, maternal age, father's education, household income, and residence	LM, Indonesia	6
Ashford, 2010 [26]	PTB	Cross-sectional	139	Number of cigarettes and biochemically verified (cotinine)	2.3 (0.96-5.96)	Mother's age, education, ethnicity; infant's gestational age; prenatal conditions	H, United States	6
Chevrier, 2008 [35]	CLP	Case-control	286	Home (y/n)	-	Region, child's sex, maternal ethnic origin	H, France	8
Crane, 2011 [17]	LBW PTB Stillbirth	Cohort	11852 11846 11852	General exposure (y/n)	- 1.87 (1.00-3.53) 3.35 (1.16-9.72)	Maternal age, nulliparity, partnered status, work status, BMI, alcohol or illicit drug use, level of education. Infant's gestational age	H, Canada	5
Fantuzzi, 2007 [36]	PTB	Case-control	1034	Number of smokers	0.92 (0.65-1.31)	Maternal age, previous PTB, hypertension, diabetes, antenatal class attendance, moderate physical activity	H, Italy	7
Gardosi, 2013 [31]	Stillbirth	Cohort	67732	General exposure (y/n)	1.3 (0.8-2.0)†	Parity, ethnicity, BMI, mental health history, diabetes; foetal growth restriction	H, United Kingdom	6
Ghosh, 2013 [18]	LBW PTB	Case-control	1024 1587	Number of people who smoked in the home	1.34 (0.84-2.16) 1.27 (0.95-1.7)	Maternal age, race, education, parity, and birthplace	H, United States	7
Goel, 2004 [37]	LBW PTB	Cohort	576 576	General exposure (y/n)	1.03 (0.65-1.65) 1.15 (0.69-1.92)	Maternal age, education, occupation, birth order, anaemia, number of live issues	LM, India	6
Guo, 2010 [38]	CLP	Case-control	400	Home or work (y/n)	-	-	UM, China	4
Hao, 2015 [34]	CLP	Case-control	842	Home or work (y/n)	2.52 (1.90, 3.33)*	-	UM, China	5

TABLE 1. Cont.

First author, year	Outcome	Study design	No. of cases/ study size	Exposure	Risk estimates (aOR where applicable)	Adjusted factors	Income level, country	QA
Honein, 2007 [39]	CLP	Case-control	4323	Home or work (y/n)	1.0 (0.8-1.3)	Folic acid supplements, infant's gender, mother's age, race/ethnicity, BMI, alcohol use, education, gravidity	H, United States	6
Hyland, 2014 [32]	Stillbirth	Cohort	38451	Home or work (y/n)	1.22 (1.02-1.47)	Mother's BMI, age, ethnicity, education, alcohol intake, parity, and ever oral contraceptive use	H, United States	7
Ion, 2015 [27]	PTB	Cohort	13799	Partner or member of household smoked	1.15 (0.81-1.63)	Maternal age, BMI, parity, ethnicity, employment status, socioeconomic status, asthma, and preeclampsia	H, United Kingdom	8
Jaddoe, 2008 [40]	LBW PTB	Cohort	7098	Number of hours/days	0.92 (0.11-7.62) 1.83 (0.63-5.30)	Maternal BMI, alcohol consumption, education, ethnicity, parity, age, and gender	H, The Netherlands	7
Jia, 2011 [41]	CLP	Case-control	216	General exposure (y/n)	9.23 (5.96-14.28)*	Infant's gender, birth weight, Maternal age, weight, multivitamin use, folic acid supplements	UM, China	5
Jianyan, 2010 [42]	CLP	Case-control	400	Home or work (y/n)	1.72 (1.08-2.74)	Maternal and paternal schooling	UM, China	6
Khader, 2011 [19]	LBW PTB	Cross-sectional	7905 8490	Home or work (y/n)	1.56 (1.31-1.89) 1.61 (1.3-1.99)	Mother's age, education, occupation, income, height, blood type, parity, history of preterm birth	UM, Jordan	6
Kharrazi, 2004 [43]	LBW PTB Stillbirth	Cohort	2777 2900 2815	Biochemical verification (cotinine)	-	-	H, United States	7
Khattar, 2013 [20]	LBW	Case-control	300	Mean number of cigarettes	3.16 (1.9-5.29)	Current preterm pregnancy, mother's socioeconomic status, previous LBW baby, severe anaemia, antenatal care	LM, India	7
Ko, 2014 [21]	LBW PTB	Cohort	21086 21086	Mean number of cigarettes	1.12 (0.88-1.42) 1.19 (0.91-1.55)	Maternal age, nationality, parity, weight gain during pregnancy, gender of infant, multi-foetus	H, Taiwan	7
Krstev, 2013 [44]	LBW	Cohort	1649	Member of household smoked	1.37 (0.73-2.57)	Newborn's sex, maternal age, residence, parity, education, employment, socioeconomic status	UM, Serbia	6

TABLE 1. Cont.

First author, year	Outcome	Study design	No. of cases/ study size	Exposure	Risk estimates (aOR where applicable)	Adjusted factors	Income level, country	QA
Li, 2010 [45]	CLP	Case-control	739	Home or work (y/n)	2.0 (1.2-3.4)	Maternal occupation, periconceptual flu or fever, and infant's sex	UM, China	6
Little, 2004 [46]	CLP	Case-control	216	Home or work (y/n)	0.9 (0.5-1.7)	Infant's sex, season of birth, ethnic group. Maternal education	H, United Kingdom	6
Luo, 2012 [28]	PTB	Case-control	722	General exposure (y/n) and biochemically verified (cotinine)	2.2 (1.56-3.12)	Mother's income, age, education level, and pre-pregnancy BMI	UM, China	9
Metgud, 2012 [22]	LBW	Cohort	1138	Partner smoked	2.3 (1.1-4.9)	Maternal education, age at 1 <sup>st</sup> pregnancy, birth interval, previous LBW baby, weight gain during pregnancy, maternal weight, hypertension, current high-risk pregnancy, antenatal registration	LM, India	3
Mojibyan, 2013 [29]	PTB	Cross-sectional	205	General exposure (y/n)	-	-	UM, Iran	3
Norsa'adah, 2014 [23]	LBW	Cohort	420	Husband or housemates smoked	1.37 (0.51-3.65)	Gestational age, maternal height, history of LBW baby, parity	UM, Malaysia	6
Ojima, 2004 [47]	LBW	Case-control	690	Home or work (y/n)	1.3 (0.88-1.93)	Maternal age, occupation, economic difficulty, previous LBW baby, medical treatment for infertility	H, Japan	5
Pogodina, 2009 [48]	LBW	Cohort	2206	General exposure (y/n)	1.29 (1.06-1.57)	Marital status and income	H, United States	6
Qiu, 2014 [30]	PTB	Cohort	11102	Home or work (y/n)	1.12 (0.95-1.32)	Maternal age, educational level, employment status, preeclampsia, diabetes, parity, history of preterm delivery, and caesarean delivery	UM, China	7
Subramoney, 2010 [24]	LBW PTB Stillbirth	Cohort	802 903 924	Partner smoked	-	-	LM, India	6
Uncu, 2005 [49]	LBW PTB	Cohort	450 450	Partner smoked	1.96 (1.15-3.34)* 2.15 (1.26-3.67)*	-	UM, Turkey	4



TABLE 1. Cont.

First author, year	Outcome	Study design	No. of cases/ study size	Exposure	Risk estimates (aOR where applicable)	Adjusted factors	Income level, country	QA
Varner, 2014 [33]	Stillbirth	Case-control	2126	Cotinine levels above 0.25 ng/ml and less than 3 ng/ml	2.10 (1.18-3.73)*	-	H, United States	7
Wahabi, 2013 [25]	LBW	Cohort	3152	General exposure (y/n)	-	-	H, Saudi Arabia	4
Ward, 2007 [50]	LBW PTB	Cohort	10347 10262	Partner smoked	1.23 (0.96-1.58) 1.21 (0.56-1.51)	Maternal age, BMI, parity, alcohol use, education, income, ethnicity, gestatio- nal diabetes	H, United Kingdom	6
Wdowiak, 2009 [51]	LBW PTB	Cohort	150 150	Home or work (y/n)	-	-	H, Poland	3
Wu, 2007 [52]	LBW PTB	Cohort	1388 1388	Partner smoked	-	-	UM, China	6

CLP – cleft lip with or without cleft palate, LBW – low birth weight, PTB – preterm birth, BMI – body mass index, aOR – adjusted odds ratio, QA – quality assessment, H – high-income, UM – upper middle-income, LM – lower middle-income. \*Crude odds ratio, †Risk ratio is presented. Jia et al. (2011) was removed from analyses due to high heterogeneity after sub-analyses were conducted.

11 for PTB [16-19, 21, 24, 26-30], five for stillbirth [17, 24, 31-33], and one for cleft lip [34], beyond those included in previous reviews. Together, the 39 studies represented 223,073 women/infants: 78,948 were included in LBW analyses; 89,831 in PTB analyses; 123,900 in stillbirth analyses; and 7422 in cleft lip analyses (some women were included in multiple analyses). SHS exposure was collected via cotinine assay in two studies [33, 43] and via self-reports and cotinine assay in two studies [26, 28]. All other SHS measures were self-reported. Two studies measured SHS exposure in the home, 12 studies measured exposure at home or at work, 11 studies measured exposure to smoking by a partner or spouse, and 14 studies did not identify the source of the SHS exposure.

Of the 39 studies, 19 (49%) were conducted in HICs, 14 were conducted in upper middle-income countries, and six were conducted in lower middle-income countries. For this review, no studies were identified that occurred in low-income countries. Study quality ratings ranged from 3.0 to 9.0, with a mean quality score of 6.0. The mean quality score for studies from HICs was 6.2; for those from lower MICs, it was 6.0; and for studies from upper MICs, it was 5.7. Twenty-nine studies (74%) were considered to be of high quality, having scores higher than the chosen threshold of 6.0. The quality of other studies was lower due to a number of factors including inadequately ascertaining SHS exposure or adjusting for confounders. The results of the meta-analyses are presented in Table 2.

## OUTCOME MEASURES

### Low birth weight

The effect of SHS exposure on the risk of low birth weight was assessed using 15 cohort/cross-sectional studies and four case-control studies in the unadjusted analyses. Fourteen studies presented aORs, three of which did not include unadjusted ORs and thus were not included in both analyses (Fig. 2). Studies were conducted in 16 countries, including three lower MICs (Gaza, India, and Indonesia) and five upper MICs (China, Jordan, Malaysia, Serbia, and Turkey); studies from these 16 countries had a mean quality assessment score of 5.8. All analyses were significant: Overall exposure to SHS was associated with a 28% increase in the risk of delivering a low birth weight infant ( $n = 12$ ; 95% CI: 1.103-1.490). Results using the adjusted high-quality sample (QA > 6) in MICs resulted in pORs higher than in the analyses of all countries but still highly significant ( $n = 7$ ; pOR: 1.351 [95% CI: 1.029-1.773]).

### Preterm birth

The effect of maternal SHS exposure on the risk of preterm birth was assessed using 14 cohort/cross-sectional studies and three case-control studies in the unadjusted analyses (Fig. 3). Thirteen studies were included in the adjusted analyses, including two studies that only

TABLE 2. Meta-analysis results

	Unadjusted full sample					Adjusted high QA sample				
	<i>n</i>	pOR	95% CI	<i>p</i> -value	<i>I</i> <sup>2</sup>	<i>n</i>	pOR	95% CI	<i>p</i> -value	<i>I</i> <sup>2</sup>
<b>Low birth weight</b>										
All countries	19	1.573	1.360-1.818	<b>0.000</b>	74.392	12	1.282	1.103-1.490	<b>0.001</b>	69.346
MICs	10	1.857	1.467-2.352	<b>0.000</b>	68.319	7	1.351	1.029-1.773	<b>0.030</b>	81.469
HICs	9	1.268	1.178-1.365	<b>0.000</b>	51.207	5	1.228	1.084-1.392	<b>0.001</b>	0.000
<b>Preterm birth</b>										
All countries	17	1.38	1.217-1.566	<b>0.000</b>	74.731	12	1.29	1.127-1.476	<b>0.000</b>	52.301
MICs	8	1.622	1.334-1.972	<b>0.000</b>	73.105	5	1.404	1.083-1.819	<b>0.010</b>	75.228
HICs	9	1.196	1.079-1.326	<b>0.001</b>	31.673	7	1.188	1.048-1.346	<b>0.007</b>	0.000
<b>Stillbirth</b>										
All countries	6	1.556	1.14-2.123	<b>0.005</b>	69.115	2	1.23	1.037-1.459	<b>0.018</b>	0.000
MICs	*	*	*	*	*	–	–	–	–	–
HICs	5	1.463	1.066-2.008	<b>0.018</b>	69.01	2	1.23	1.037-1.459	<b>0.018</b>	0.000
<b>Orofacial clefts</b>										
All countries	8	2.045	11.277-3.273	<b>0.003</b>	92.402	4	1.307	0.893-1.92	0.168	67.328
1 study removed	7	1.657	1.201-2.287	<b>0.002</b>	81.627	4	1.307	0.893-1.92	0.168	67.328
MICs	5	2.785	1.637-4.740	<b>0.000</b>	89.377	2	1.839	1.3-2.602	<b>0.001</b>	0.000
1 study removed	4	2.181	1.816-2.619	<b>0.000</b>	0.000	2	1.839	1.3-2.602	<b>0.001</b>	0.000
HICs	3	1.166	0.872-1.561	0.301	40.849	2	0.987	0.795-1.225	0.905	0.000

pOR – pooled odds ratio, CI – confidence interval, MIC – middle-income country, HIC – high-income country. \*one study fit criteria, no meta-analysis conducted. Bold *p*-values are significant at the level *p* = 0.05.

presented aORs. Studies were conducted in 19 countries, including two lower MICs (India and Indonesia) and four upper MICs (China, Iran, Jordan, and Turkey); studies in these 19 countries had a mean QA score of 6.1. All analyses were significant. Overall exposure to SHS was associated with a 29% increase in the risk of delivering an infant preterm (*n* = 12; 95% CI: 1.127-1.476). Results using a high quality-adjusted (QA > 6) sample in MICs resulted in pORs higher than in the analyses of all countries and still highly significant (*n* = 7; pOR: 1.404 [95% CI: 1.083-1.819]).

#### Stillbirth

The effect of SHS exposure on the risk of stillbirth was assessed in five cohort/cross-sectional studies and one case-control study (Fig. 4). Three of the six studies were also included in adjusted analyses. Studies were conducted in four countries, including one lower middle-income country (India), and had a mean QA score of 6.3. Due to only one study in an MIC, analyses could not be stratified by country income group. All results were significant except for the sub-analysis using only adjusted data. Using adjusted high-quality studies, SHS exposure was associated with a 23% increase in the risk of having a stillbirth (*n* = 2; 95% CI: 1.037-1.459).

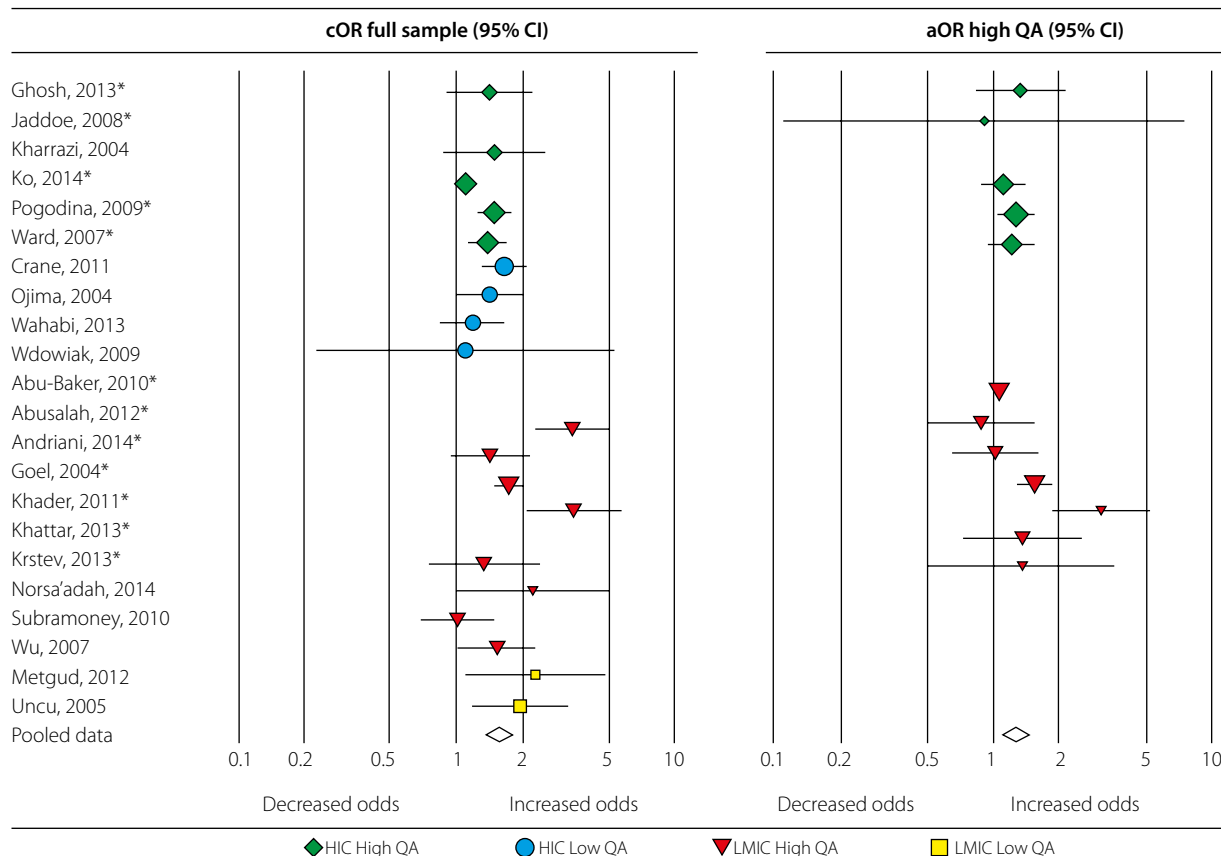
#### Orofacial clefts

Eight studies were included in the unadjusted CLP analyses, of which four studies were included in the adjusted analyses, all of which were case-controls. Three studies were conducted in HICs, and five were conducted in China (an upper middle-income country). These eight studies had a mean QA score of 5.75. *I*<sup>2</sup> values were extremely high for this analysis. By removing individual studies from the analysis, it was determined that the Jia *et al.* [41] study was driving the heterogeneity; for this reason, it was removed from analyses of the full sample, and due to low QA it was not included in the statistics presented below or in Figure 5. Analyses of high-quality studies show that SHS exposure was associated with a 31% increase in the odds of CLP (*n* = 4, 95% CI: 0.893-1.92) (Fig. 5). Studies from MICs alone show that exposure to SHS is associated with an 84% increased risk (*n* = 2; 95% CI: 1.30-2.602).

#### DISCUSSION

Consistent with previous meta-analyses, a significant increase in odds of LBW, PTB, and stillbirth associated with SHS exposure was found in both the full sample and in studies from HICs and MICs. Unlike previous meta-analyses, the current work included studies





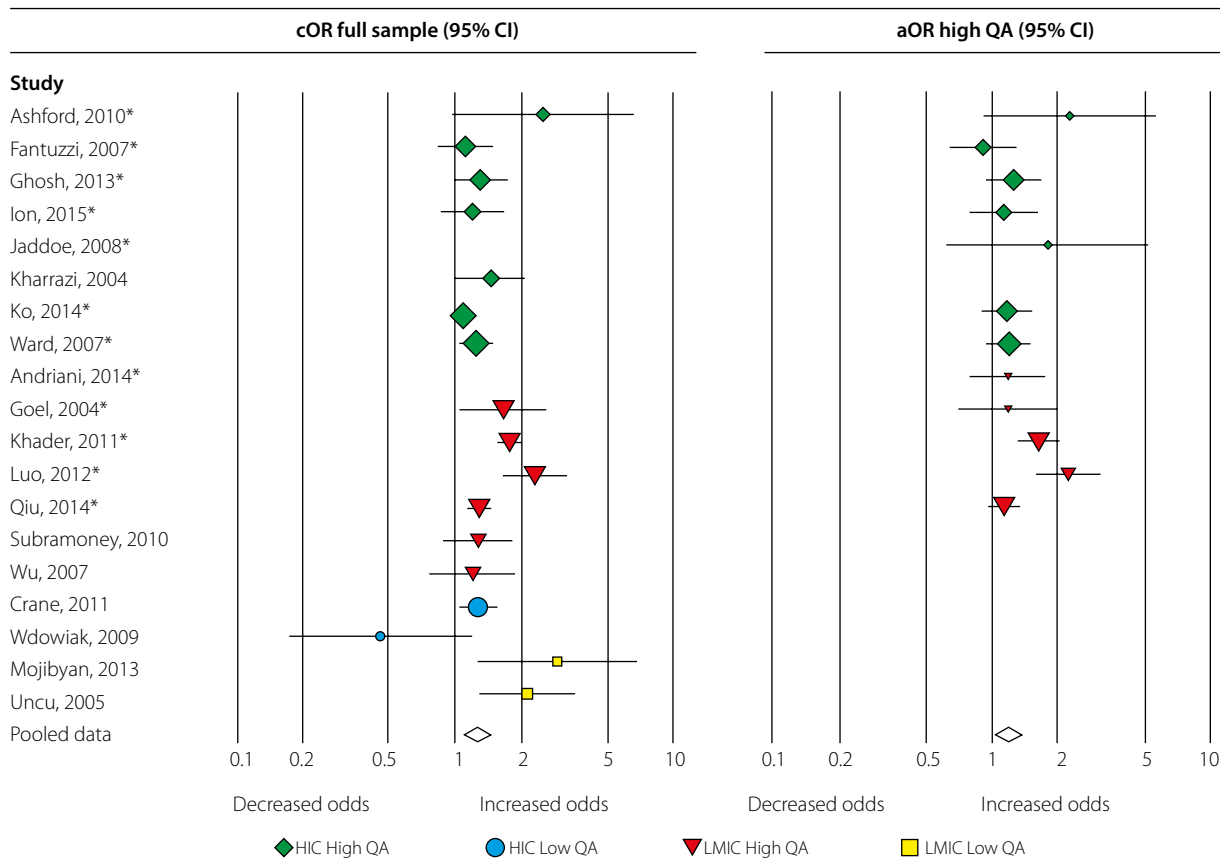
cOR – crude odds ratio, aOR – adjusted odds ratio, \*adjusted odds ratios included in high-quality analysis, CI – confidence interval, QA – quality assessment, HIC – high-income country, LMIC – low- or middle-income country

**FIG. 2.** Low birth weight forest plot

from MICs. For LBW, previous meta-analyses found pORs between 1.17 and 1.60. In the current meta-analysis, findings for both the unadjusted full sample (pOR = 1.573) and the adjusted high-quality sample (pOR = 1.282) fall within this range. The findings for stillbirth among the adjusted high-quality studies in this analysis (pOR = 1.23) were comparable with previous meta-analyses; however, this analysis was limited to HICs because of the lack of studies assessing stillbirth from LMICs. The odds of PTB among the high-quality studies were higher in our analysis (pOR = 1.29 across all studies; pOR = 1.4 in MICs) than in previous meta-analyses (pOR = 1.16-1.18). The current meta-analysis also found that odds for CLP using the unadjusted data (pOR = 2.045) were consistent with previous reports (pOR = 2.05) [11] but were much lower and not statistically significant when using the adjusted high-quality sample (pOR = 1.307). The current meta-analysis found that studies in the MIC group resulted in higher pORs than those for HICs (for example, 1.4 for preterm birth in MICs vs. 1.19 in HICs). This trend holds true for all sub-analyses conducted. These differences may be due to greater exposure levels

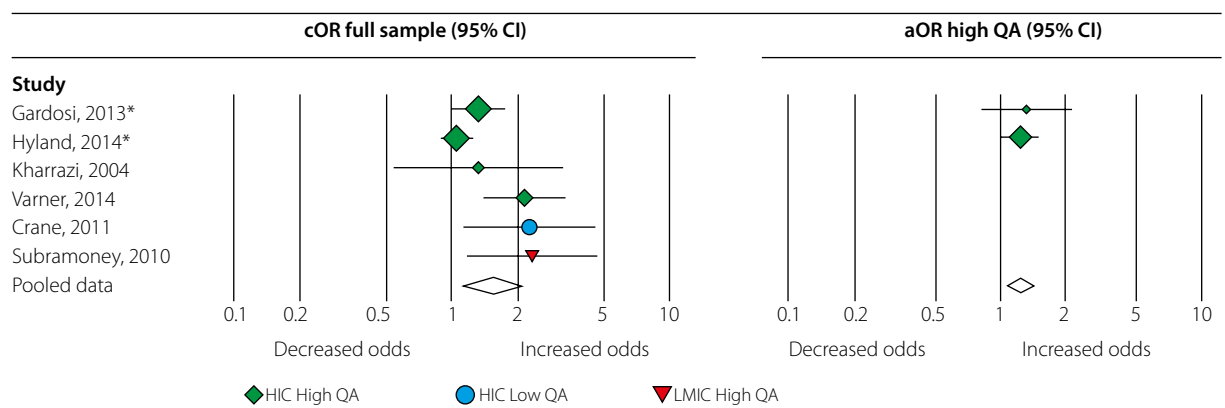
and weaker protections from SHS in MICs, differences in smoked products and patterns of use, or possible interactions with other reproductive health risks. However, given that there are still relatively few high-quality studies from LMICs, future studies should further explore these differences.

Substantial uncertainty remains around the extent of the burden on maternal and child health from SHS exposure, particularly in low- and middle-income countries. One survey of nine developing countries reported that SHS exposure during pregnancy ranged widely, from 17.1% in the Democratic Republic of Congo to 91.6% in Pakistan, but comprehensive data on SHS exposure during pregnancy is lacking from most LMICs [53,54]. Although data is limited, estimates suggest that more than one-third of nonsmoking women (35%) of reproductive age could be exposed to SHS in LMICs [55]. Thus, among women and children in many LMICs, the disease burden from SHS exposure may equal to or exceed that from direct smoking. Additionally, within LMICs, exposure to SHS may be even greater among groups with lower socioeconomic status, because low educational attainment and low wealth



cOR – crude odds ratio, aOR – adjusted odds ratio, \*adjusted odds ratios included in high-quality analysis, CI – confidence interval, QA – quality assessment, HIC – high-income country, LMIC – low- or middle-income country

FIG. 3. Preterm birth forest plot

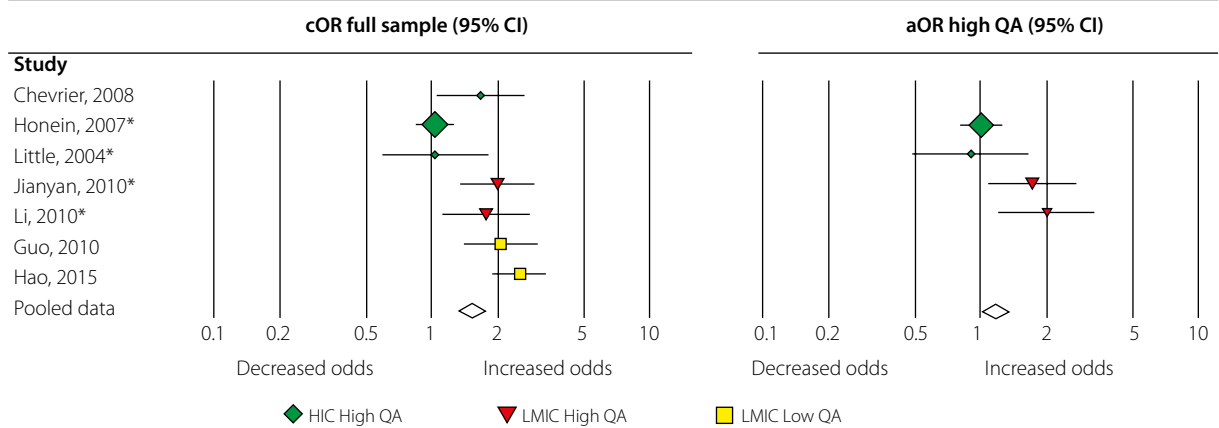


cOR – crude odds ratio, aOR – adjusted odds ratio, \*adjusted odds ratios included in high-quality analysis, CI – confidence interval, QA – quality assessment, HIC – high-income country, LMIC – low- or middle-income country

FIG. 4. Stillbirth forest plot

tend to be associated with increased daily SHS exposure [56]. The relative risks reported for adverse reproductive outcomes among mothers who smoke during pregnancy are not dramatically different from the relative risks found here for SHS exposure. Studies of maternal smok-

ing associated with low-birth weight and orofacial clefts have reported odds ratios up to 2.5, which is higher than the effects found in this analysis, but results are similar for PTB and stillbirth. Thus, it is reasonable to expect that the far greater prevalence of SHS exposure during pregnan-



cOR – crude odds ratio, aOR – adjusted odds ratio, \*adjusted odds ratios included in high-quality analysis, CI – confidence interval, QA – quality assessment, HIC – high-income country, LMIC – low- or middle-income country

**FIG. 5.** Cleft lip with or without cleft palate forest plot

cy may have a greater impact, particularly in those countries where smoking is very low among women but high among men [57].

In 2013, the World Health Organisation published recommendations for the prevention and management of tobacco use and SHS exposure in pregnancy, noting that pregnancy is often a window of opportunity to intervene with women and their families [58]. However, evidence on effective interventions is limited, especially in LMICs. A systematic review found that clinical interventions delivered in prenatal care settings appear to reduce SHS exposure, but that more rigorous studies, using biochemical validation, are needed [59]. A multi-method study from Indonesia found high levels of exposure to SHS as well as high levels of awareness about risks to children, but very low prevalence of household indoor smoking rules. While women expressed a low sense of self-efficacy about persuading their husbands to quit smoking in their homes, community-wide smoking bans had substantial support [60]. Evidence from HICs has shown that smoke-free policies have led to reductions in adverse maternal and child health outcomes [61, 62]. Additionally, a recent study in 23 European Union countries found that increases in the median price of cigarettes were associated with reductions in infant mortality between 2004 and 2014, probably because of reductions in smoking and secondhand smoke exposure [63]. However, more research is needed to develop and evaluate interventions to address SHS exposure during pregnancy in LMICs in the context of cultural and socioeconomic barriers [4].

Limitations of the current study include the high heterogeneity ( $I^2 > 75\%$ ) of some of the analyses, which may suggest a lack of generalisability. Because high heterogeneity was more common in the MIC analyses than the HIC analyses, it may reflect differences in smoking

prevalence and smoking intensity across MICs; it could also reflect differences in exposure assessment. Notably, heterogeneity was high in all studies on cleft palate in MICs although all five studies were conducted in China. It is possible that the relatively low quality of these five studies impacted the risk estimates for cleft palate.

Additional limitations relate to the measurement of reproductive outcomes and SHS exposure in the studies used for this analysis. For outcomes of preterm birth, almost all studies from MICs only used the last menstrual date to determine whether the birth was preterm, and most studies from MICs did not confirm gestational age when looking at LBW. Preterm birth and stillbirth can be difficult to confirm in low-resource settings because of the difficulty of accurately assessing gestational age. Collecting data about birth outcomes via birth records also constitutes another limitation of this meta-analysis. Birth outcomes may not be reported if the birth does not occur in a clinic, and high-risk pregnancies are more often referred to clinics for births in some settings, so the rates may be overestimated for certain outcomes. Lastly, in most studies SHS exposure during pregnancy was estimated by self-reported household smoking status without biochemical verification. It is especially challenging to precisely assess SHS exposure from self-report because exposure may be influenced not only by frequency of smoking but also by conditions in the home and the amount of time spent in the home.

This study demonstrates that SHS exposure poses a serious threat for maternal and child health in LMICs as well as HICs. In fact, the magnitude of the risk for birth outcomes included in the analysis was greater among the MIC studies than the HIC studies. More high-quality studies of SHS exposure during pregnancy are needed from LMICs to better characterise the magnitude of the risk and understand the factors behind these

differences. Additionally, more comprehensive monitoring of SHS exposure during pregnancy and evaluation of potential interventions are needed to effectively address the adverse impact of SHS exposure during pregnancy in parts of the world where the burden is greatest.

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## DISCLOSURE

The authors report no conflict of interest.

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#### AUTHOR'S CONTRIBUTIONS

MP, ALK, and JB prepared the research concept and design, collected data, wrote the article, revised the article and finally approved it.